## Statistical Inference

## Test Set 2

1. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from a $U(-\theta, 2 \theta)$ population. Find MLE of $\theta$.
2. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from a Pareto population with density $f_{X}(x)=\frac{\beta \alpha^{\beta}}{x^{\beta+1}}, x>\alpha, \alpha>0, \beta>2$. Find the MLEs of $\alpha, \beta$.
3. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from a $U(-\theta, \theta)$ population. Find the MLE of $\theta$.
4. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from a lognormal population with density $f_{X}(x)=\frac{1}{\sigma x \sqrt{2 \pi}} \exp \left\{-\frac{1}{2 \sigma^{2}}\left(\log _{e} x-\mu\right)^{2}\right\}, x>0$. Find the MLEs of $\mu$ and $\sigma^{2}$.
5. Let $\left(X_{1}, Y_{1}\right),\left(X_{2}, Y_{2}\right), \ldots,\left(X_{n}, Y_{n}\right)$ be a random sample from a bivariate normal population with parameters $\mu_{1}, \mu_{2}, \sigma_{1}^{2}, \sigma_{2}^{2}, \rho$. Find the MLEs of parameters.
6. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from an inverse Gaussian distribution with density $f_{X}(x)=\left(\frac{\lambda}{2 \pi x^{3}}\right)^{1 / 2} \exp \left\{-\frac{\lambda(x-\mu)^{2}}{2 \mu^{2} x}\right\}, x>0$. Find the MLEs of parameters.
7. Let $\left(X_{1}, X_{2}, \ldots, X_{k}\right)$ have a multinomial distribution with parameters $n=\sum_{i=1}^{k} X_{i}$, $p_{1}, \ldots, p_{k} ; 0 \leq p_{1}, \ldots, p_{k} \leq 1, \sum_{j=1}^{k} p_{j}=1$, where $n$ is known. Find the MLEs of $p_{1}, \ldots, p_{k}$.
8. Let one observation be taken on a discrete random variable $X$ with pmf $p(x \mid \theta)$, given below, where $\Theta=\{1,2,3\}$ Find the MLE of $\theta$.

|  |  | $\theta$ |  |  |
| :---: | :---: | :---: | :---: | :---: |
| x |  | 1 | 2 | 3 |
|  | 1 | $1 / 2$ | $1 / 4$ | $1 / 4$ |
|  | 2 | $3 / 5$ | $1 / 5$ | $1 / 5$ |
|  | 3 | $1 / 3$ | $1 / 2$ | $1 / 6$ |
|  | 4 | $1 / 6$ | $1 / 6$ | $2 / 3$ |

9. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from the truncated double exponential distribution with the density

$$
f_{X}(x)=\frac{e^{-|x|}}{2\left(1-e^{-\theta}\right)},|x|<\theta, \theta>0 .
$$

Find the MLE of $\theta$.
10. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from the Weibull distribution with the density $f_{X}(x)=\alpha \beta x^{\beta-1} e^{-\alpha x^{\beta}}, x>0, \alpha>0, \beta>0$.
Find MLE of $\alpha$ when $\beta$ is known.

## Hints and Solutions

1. The likelihood function is $L(\theta, \underline{x})=\frac{1}{(3 \theta)^{n}},-\theta<x_{(1)} \leq x_{(2)} \leq \cdots \leq x_{(n)}<2 \theta, \theta>0$. Clearly it is maximized with respect to $\theta$, when $\theta$ takes its infimum. Hence, $\hat{\theta}_{M L}=\max \left(-X_{(1)}, \frac{X_{(n)}}{2}\right)$.
2. The likelihood function is $L(\alpha, \beta, \underline{x})=\frac{\beta^{n} \alpha^{n \beta}}{\left(\prod_{i=1}^{n} x_{i}\right)^{\beta+1}}, x_{(1)}>\alpha, \alpha>0, \beta>2 . L$ is maximized with respect to $\alpha$ when $\alpha$ takes its maximum. Hence $\hat{\alpha}_{M L}=X_{(1)}$. Using this we can rewrite the likelihood function as $L^{\prime}(\beta, \underline{x})=\frac{\beta^{n}\left\{x_{(1)}\right\}^{n \beta}}{\left(\prod_{i=1}^{n} x_{i}\right)^{\beta+1}}, \beta>2$. The log likelihood is $\log L^{\prime}(\beta, \underline{x})=n \log \beta+n \beta \log x_{(1)}-(\beta+1) \log \left(\prod_{i=1}^{n} x_{i}\right)$. This can be easily maximized with respect to $\beta$ and we get $\hat{\beta}_{M L}=\left[\frac{1}{n} \sum \log \frac{X_{(i)}}{X_{(1)}}\right]^{-1}$.
3. Arguing as in Sol. 1, we get $\hat{\theta}_{M L}=\max \left(-X_{(1)}, X_{(n)}\right)=\max _{1 \leq i \leq n}\left|X_{i}\right|$.
4. Directly maximizing the log-likelihood function with respect to $\mu$ and $\sigma^{2}$, we get

$$
\hat{\mu}_{M L}=\frac{1}{n} \sum \log X_{i}, \hat{\sigma}_{M L}^{2}=\frac{1}{n} \sum\left(\log X_{i}-\hat{\mu}_{M L}\right)^{2} .
$$

5. The maximum likelihood estimators are given by

$$
\hat{\mu}_{1}=\bar{X}, \hat{\mu}_{2}=\bar{Y}, \hat{\sigma}_{1}^{2}=\frac{1}{n} \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)^{2}, \hat{\sigma}_{2}^{2}=\frac{1}{n} \sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}, \hat{\rho}=\frac{1}{n} \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)\left(Y_{i}-\bar{Y}\right) /\left(\hat{\sigma}_{1} \hat{\sigma}_{2}\right) .
$$

6. The maximum likelihood estimators are given by

$$
\hat{\mu}_{M L}=\bar{X}, \hat{\lambda}_{M L}=\left[\frac{1}{n}\left(\sum_{i=1}^{n} \frac{1}{X_{i}}-\frac{1}{\bar{X}}\right)\right]^{-1}
$$

7. The maximum likelihood estimators are given by

$$
\hat{p}_{1}=\frac{X_{1}}{n}, \cdots, \hat{p}_{k}=\frac{X_{k}}{n}
$$

8. $\hat{\theta}_{M L}=1$, if $x=1,2$

$$
\begin{array}{ll}
=2, & \text { if } x=3 \\
=3, & \text { if } x=4
\end{array}
$$

9. $\hat{\theta}_{M L}=\max \left(-X_{(1)}, X_{(n)}\right)=\max _{1 \leq i \leq n}\left|X_{i}\right|$.
10. $\hat{\alpha}_{M L}=\frac{n}{\sum x_{i}^{\beta}}$
